

(5)

Theorem

$Y_1, \dots, Y_n \sim \text{i.i.d. } f(y; \theta)$ which satisfies the regularity assn. And. $I(\theta)$ is cont

$$TLR = -2 \{ \ell(\tilde{\theta}) - \ell(\hat{\theta}) \}$$

Under $H_0: \theta_1 = \theta_{10}$, $TLR \xrightarrow{d} \chi^2_r$ as $n \rightarrow \infty$

proof

$$\ell(\tilde{\theta}) = \ell(\hat{\theta}) + \underbrace{(\tilde{\theta} - \hat{\theta})'}_{=0} \ell_{\theta}(\hat{\theta}) + \frac{1}{2} (\tilde{\theta} - \hat{\theta})' \ell_{\theta\theta}(\hat{\theta}^*) (\tilde{\theta} - \hat{\theta})$$

$$-2 \{ \ell(\tilde{\theta}) - \ell(\hat{\theta}) \} = (\tilde{\theta} - \hat{\theta})' \left\{ -\ell_{\theta\theta}(\hat{\theta}^*) \right\} (\tilde{\theta} - \hat{\theta})$$

Road map. 1) find asy dist'n of $(\tilde{\theta} - \hat{\theta}) \sqrt{n}$

2) find limit in probs of $-\ell_{\theta\theta}(\hat{\theta}^*) \frac{1}{n}$

We start with 2) $\ell_{\theta\theta}(\hat{\theta}^*) = \ell_{\theta\theta}(\theta_0) + \sum_{j=1}^r \frac{\partial^2 \ell(\hat{\theta}^{**})}{\partial \theta_1 \partial \theta_1 \partial \theta_j} (\theta_j^* - \theta_j)$

$$\left| -\frac{1}{n} \ell_{\theta\theta}(\hat{\theta}^*) + \frac{1}{n} \ell_{\theta\theta}(\theta_0) \right| \leq \frac{1}{n} \sum_{j=1}^r h(Y_{1j}) \cdot (\theta_j^* - \theta_j)$$

$$\Rightarrow -\frac{1}{n} \ell_{\theta\theta}(\hat{\theta}^*) \xrightarrow{P} I(\theta_0) \quad \begin{matrix} O_p(1) & O_p(1) \end{matrix}$$

since $-\frac{1}{n} \ell_{\theta\theta}(\theta_0) \xrightarrow{P} I(\theta_0)$

To prove 1) ; recall

$$\hat{\theta} - \theta_0 = \frac{1}{n} \sum I^{-1}(\theta_0) \frac{d}{d\theta} \log f(Y_i; \theta_0) + R_n$$

$$\tilde{\theta} - \theta_{20} = \frac{1}{n} \sum I_{22}^{-1}(\theta_0) \frac{d}{d\theta_2} \log f(Y_i; \theta_0) + Q_n$$

Now, we want $\tilde{\theta} - \theta_0$; for this we make up a matrix $H(\theta)$ by padding $I_{22}^{-1}(\theta_0)$ with 0-block matrices

$$H(\theta) \stackrel{\text{def}}{=} \begin{pmatrix} \mathbf{0} & \mathbf{0} \\ \mathbf{0} & I_{22}^{-1}(\theta_0) \end{pmatrix} ; \tilde{Q}_n = \begin{pmatrix} \mathbf{0} \\ Q_n \end{pmatrix}$$

and obtain

$$\tilde{Q}_n = o_p(n^{1/2})$$

$$\tilde{\theta} - \theta_0 = \frac{1}{n} \sum_{i=1}^n H(\theta_0) \frac{d}{d\theta} \log f(Y_i; \theta) + \tilde{Q}_n$$

$$\tilde{\theta} - \hat{\theta} = \frac{1}{n} \sum_{i=1}^n \{ H(\theta_0) - I^{-1}(\theta_0) \} \frac{d}{d\theta} \log f(Y_i; \theta_0) + o_p(n^{-1/2})$$

$$\begin{aligned} \text{cov} \left(\left\{ H(\theta_0) - I^{-1}(\theta_0) \right\} \frac{d}{d\theta} \log f(Y_i; \theta_0) \right) \\ = \left\{ H(\theta_0) - I^{-1}(\theta_0) \right\} \cdot I(\theta_0) \left\{ H(\theta_0) - I^{-1}(\theta_0) \right\}^T \\ = H(\theta_0) I(\theta_0)^T H(\theta_0) - \underbrace{H(\theta_0)}_{=H(\theta_0)} - H(\theta_0) + I^{-1}(\theta_0) = I^{-1}(\theta_0) - H(\theta_0) \end{aligned}$$

$$\Rightarrow \sqrt{n}(\tilde{\theta} - \hat{\theta}) \xrightarrow{d} N_b \left(0, I^{-1}(\theta_0) - H(\theta_0) \right)$$

Note the asymptotic variance is of dim $b \times b$. $I^{-1}(\theta_0) \left\{ I_{bb} - I(\theta_0) H(\theta_0) \right\}$
 However we have result $= I^{-1}(\theta_0) \begin{pmatrix} I_{rr} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} \end{pmatrix}$

$$\left[\begin{array}{l} X_n \xrightarrow{d} N(0, \Sigma) \text{ and } D_n \xrightarrow{P} D \text{ then } \\ \text{if } D\Sigma \text{ is idempotent and } r = \text{rank } D\Sigma = \text{Trace } D\Sigma \\ X_n^T D_n X_n \xrightarrow{d} \chi_r^2 \end{array} \right.$$

Use this result for $X_n = \sqrt{n}(\tilde{\theta} - \hat{\theta})$; $\Sigma = I^{-1}(\theta_0) - H(\theta_0)$; $D = I(\theta_0)$
 after checking that $D\Sigma = I_{b \times b} - I(\theta_0) H(\theta_0)$ which is idemp and
 $\text{Trace } D\Sigma = b - (b-r) = r \Rightarrow n(\tilde{\theta} - \hat{\theta})^T \{ \dots \} (\tilde{\theta} - \hat{\theta}) \xrightarrow{d} \chi_r^2$